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Research Article

An Improvement in Variance Estimator for the Estimation of Population Variance, Using Known Values of Auxiliary Information

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ABSTRACT

In this research paper, we have developed an efficient estimator by introducing the linear combination of deciles arithmetic mean and coefficient of kurtosis as auxiliary information to achieve efficiency of the estimator for the estimation of population variance. The bias and mean square error has been derived up to the first order of approximation. The empirical study has been carried out through numerical demonstration to justify the performance of new developed estimators.

Key words: Deciles arithmetic mean, Coefficient of kurtosis, Bias, Mean Square Error, Efficiency.

INTRODUCTION

In survey sampling, the functions of population parameters of auxiliary variable have great role in modifying and developing the new estimators when there exists a close association between auxiliary and study variables. Different authors have utilized this auxiliary information in different ways in order to increase the efficiency of estimators. Some of them from literature who have addressed the issue of utilizing supplementary information are Isaki1 who has utilized standard deviation as auxiliary information to enhance the efficiency of the estimator for the estimation of population variance. Upadhyaya and Singh² have incorporated the coefficient of kurtosis as auxiliary variable to increase the efficiency of estimator for the estimation of population variance. Kadilar and Cingi³ introduced the coefficient of variation as

auxiliary variable to improve the efficiency of the estimator. On the same lines, authors such as Sarandal CE^4 , M.A. Bhat⁵ have also utilized this auxiliary information to enhance the efficiency of estimators The Strategy of modifying estimators, by using auxiliary information is now being regularly carried out in the field of survey sampling to improve the efficiency of estimators in order to have the precise and reliable estimates in survey estimation.

Let us consider a finite population having Ndistinct and identifiable units. Let $Y_i (i = 1, 2, 3, ..., N)$ denote the observations on study variable Y and $X_i (i = 1, 2, 3, ..., N)$ denote the observations on auxiliary variable (X_i), when the information about auxiliary variable is known.

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In this Study our aim is to estimate the finite population variance by introducing new and improved estimators in sample survey's.

MATERIAL AND METHODS

Notations

 $N = population \quad size, \ n = Sample \quad size, \ \gamma = \frac{1}{n}, Y = study \quad var \ iable, X = Auxiliary \quad var \ iable,$ $\overline{X} & \overline{Y} = Population \quad means, \overline{x} & \overline{y} = Sample \quad means, S_y^2 & S_x^2 = Population \quad var \ iances,$ $C_x & C_y = Cofficient \quad of \quad var \ iations, \ \rho = correlation \quad cofficient, \quad \beta_{1x} = Sekewness & \\ \beta_{2x} = Kurtosis \\ D = Decile, D_{A,M} = Decile \quad airthmetic \quad mean, \ B(.) = Bias & MSE(.) = Mean \quad square \ error \\ Proposed Estimator: \quad \hat{S}_M^2 = s_y^2 \left[\frac{S_x^2 + (D_{A,M}, \beta_{2x})}{s_x^2 + (D_{A,M}, \beta_{2x})} \right]$

We have derived the bias and mean square error of the proposed estimator up to the first order of approximation as given below:

Let
$$e_0 = \frac{s_y^2 - S_y^2}{S_y^2}$$
 and $e_1 = \frac{s_x^2 - S_x^2}{S_x^2}$. Further we can write $s_y^2 = S_y^2(1 + e_0)$ and $s_x^2 = S_x^2(1 + e_1)$ and

from the definition of e_0 and e_1 we obtain:

$$E[e_0] = E[e_1] = 0, \qquad E[e_0^2] = \frac{1-f}{n} (\beta_{2(y)} - 1), \qquad E[e_1^2] = \frac{1-f}{n} (\beta_{2(x)} - 1),$$
$$E[e_0e_1] = \frac{1-f}{n} (\lambda_{22} - 1)$$
The proposed estimator is given as:

$$\hat{S}_{M}^{2} = s_{y}^{2} \left[\frac{S_{x}^{2} + \alpha a_{i}}{s_{x}^{2} + \alpha a_{i}} \right] \qquad (4.1)$$

$$\Rightarrow \quad \hat{S}_{M}^{2} = s_{y}^{2} (1 + e_{0}) \left[\frac{S_{x}^{2} + \alpha a_{i}}{s_{x}^{2} + e_{1} S_{x}^{2} + \alpha a_{i}} \right]$$

$$\Rightarrow \hat{S}_{M}^{2} = \frac{S_{y}^{2}(1+e_{0})}{(1+A_{M}e_{1})} \text{ Where } A_{M} = \frac{S_{x}^{2}}{S_{x}^{2}+\alpha a_{i}}$$
$$a_{i} = (D_{AM}.\beta_{2x}) and\alpha = 1$$

$$\Rightarrow \quad \hat{S}_{M}^{2} = S_{y}^{2} (1 + e_{0}) (1 + A_{M} e_{1})^{-1}$$
(4.2)

$$\Rightarrow \quad \hat{S}_{M}^{2} = S_{y}^{2} (1 + e_{0}) (1 - A_{M} e_{1} + A_{M}^{2} e_{1}^{2} - A_{M}^{3} e_{1}^{3} + \dots)$$
(4.3.)

Expanding and neglecting the terms more than 3rd order, we get

$$\hat{S}_{M}^{2} = S_{y}^{2} + S_{y}^{2}e_{0} - S_{y}^{2}A_{M}e_{1} - S_{y}^{2}A_{M}e_{0}e_{1} + S_{y}^{2}A_{M}^{2}e_{1}^{2}$$

$$(4.5)$$

$$\Rightarrow \hat{S}_{M}^{2} - S_{y}^{2} = S_{y}^{2} e_{0} - S_{y}^{2} A_{M} e_{1} - S_{y}^{2} A_{M} e_{0} e_{1} + S_{y}^{2} A_{M}^{2} e_{1}^{2}$$
(4.6)

By taking expectation on both sides of (4.6), we get

$$E(\hat{S}_{M}^{2} - S_{y}^{2}) = S_{y}^{2}E(e_{0}) - S_{y}^{2}A_{M}E(e_{1}) - S_{y}^{2}A_{M}E(e_{0}e_{1}) + S_{y}^{2}A_{M}^{2}E(e_{1}^{2})$$

$$(4.7)$$

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$$Bias(\hat{S}_{M}^{2}) = S_{y}^{2}A_{M}^{2}E(e_{1}^{2}) - S_{y}^{2}A_{M}E(e_{0}e_{1})$$

$$Bias(\hat{S}_{M}^{2}) = \gamma S_{y}^{2}A_{M}[A_{M}(\beta_{2x}-1) - (\lambda_{22}-1)]$$
(4.8)
(4.9)

Squaring both sides of (4.7) and neglecting the terms more than 2^{nd} order and taking expectation, we get

$$E(\hat{S}_{M}^{2} - S_{y})^{2} = \gamma S_{y}^{4} E(e_{0}^{2}) + S_{y}^{4} A_{M}^{2} E(e_{1}^{2}) - 2S_{y}^{4} A_{M} E(e_{0}e_{1})$$
$$MSE(\hat{S}_{M}^{2}) = \gamma S_{y}^{4} [(\beta_{2y} - 1) + A_{M}^{2} (\beta_{2x} - 1) - 2A_{M} (\lambda_{22} - 1)]$$

Efficiency conditions: - we have derived the efficiency conditions of proposed estimators with other existing estimators under which proposed estimators are performing better than the existing estimators

The bias and Mean square error of existing ratio type estimators up to the first order of approximation is given by

1)

$$Bias(\hat{S}_{K}^{2}) = \gamma S_{y}^{2} R_{K} [R_{K}(\beta_{2x}-1) - (\lambda_{22}-1)]_{(5.1)}$$

$$MSE(\hat{S}_{K}^{2}) = \gamma S_{y}^{4} [(\beta_{2y}-1) + R_{K}^{2}(\beta_{2x}-1) - 2R_{K}(\lambda_{22}-1)]$$

$$R_{K} = Existing \ .cons \ tan \ t$$
(5.2)

, *K* = 1,2,3,4.....

Bias, MSE and constant of proposed estimators is given by

$$Bias(\hat{S}_{p}^{2}) = \gamma S_{y}^{2} R_{p} [R_{p} (\beta_{2x} - 1) - (\lambda_{22} - 1)]$$
(5.3)

$$MSE(\hat{S}_{P}^{2}) = \gamma S_{y}^{4} [(\beta_{2y} - 1) + R_{P}^{2}(\beta_{2x} - 1) - 2R_{P}(\lambda_{22} - 1)]$$
(5.4)

 $R_p = proposed.cons \tan t$ P = 1,2,3....

From Equation (2) and (3), we have $(\mathbf{P} + \mathbf{P}) \setminus \mathbf{Q}$

$$MSE(\hat{S}_{P}^{2}) \leq MSE(\hat{S}_{K}^{2}) f\lambda_{22} \geq 1 + \frac{(R_{P} + R_{K})(\beta_{2x} - 1)}{2}$$

$$MSE(\hat{S}_{P}^{2}) \leq MSE(\hat{S}_{K}^{2})$$

$$\gamma S_{y}^{4}[(\beta_{2y} - 1) + R_{P}^{2}(\beta_{2x} - 1) - 2R_{P}(\lambda_{22} - 1)] \leq \gamma S_{y}^{4}[(\beta_{2y} - 1) + R_{K}^{2}(\beta_{2x} - 1) - 2R_{K}(\lambda_{22} - 1)]$$

$$(5.5)$$

$$\Rightarrow [(\beta_{2y} - 1) + R_{P}^{2}(\beta_{2x} - 1) - 2R_{P}(\lambda_{22} - 1)] \leq [(\beta_{2y} - 1) + R_{K}^{2}(\beta_{2x} - 1) - 2R_{K}(\lambda_{22} - 1)] \quad (5.6)$$

$$\Rightarrow [1 + R_{P}^{2}(\beta_{2x} - 1) - 2R_{P}(\lambda_{22} - 1)] \leq [1 + R_{K}^{2}(\beta_{2x} - 1) - 2R_{K}(\lambda_{22} - 1)] \quad (5.7)$$

$$\Rightarrow (\beta_{2x} - 1)(R_{P}^{2} - R_{K}^{2}) [-2R_{P}(\lambda_{22} - 1)] \leq [-2R_{K}(\lambda_{22} - 1)] \quad (5.8)$$

$$\Rightarrow (\beta_{2x} - 1)(R_{P}^{2} - R_{K}^{2}) [-2(\lambda - 1)(R_{P} - R_{P})] \leq 0 \quad (5.9)$$

$$\Rightarrow (\beta_{2x} - 1)(R_P^2 - R_K^2) [-2(\lambda_{22} - 1)(R_P - R_K)] \le 0$$

$$\Rightarrow (\beta_{2x} - 1)(R_P^2 - R_K^2) \le [2(\lambda_{22} - 1)(R_P - R_K)]$$
(5.9)
(5.10)

$$\Rightarrow \left(\beta_{2x} - 1\right) \le \frac{2(\lambda_{22} - 1)(R_P - R_K)}{\left(R_P^2 - R_K^2\right)}$$
(5.11)

$$\Rightarrow \left(\beta_{2x} - 1\right) \le \frac{2(\lambda_{22} - 1)(R_P - R_K)}{(R_P - R_K)(R_P + R_K)}$$

$$(5.12)$$

$$\Rightarrow \left(\beta_{2x} - 1\right) \left(R_P + R_K\right) \le 2\left(\lambda_{22} - 1\right) \tag{5.13}$$

By solving equation (5.13), we get

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$$MSE\left(\hat{S}_{P}^{2}\right) \leq MSE\left(\hat{S}_{K}^{2}\right) f\lambda_{22} \geq 1 + \frac{\left(R_{P} + R_{K}\right)\left(\beta_{2x} - 1\right)}{2}$$

Numerical illustration:-

$$\begin{split} N &= 234, n = 35, \overline{Y} = 29.3626, \overline{X} = 245.088, S_y = 51.556, S_x = 596.332, \rho = 0.96, \beta_{2y} = 89.231, \beta_{2x} = 89.189\\ \lambda_{22} &= 4.041, \beta_{1x} = 8.83, \beta_{1y} = 8.27, TM = 167.4, Q_1 = 67.75, Q_2 = 113.5, Q_3 = 230.25, C_x = 2.43, D_1 = 49.0, D_2 = 63.0, D_3 = 75.0, D_4 = 90.0, D_5 = 113.5, D_6 = 145.9, D_7 = 197.9, D_8 = 271.1, D_9 = 467.5, D_{10} = 6720.0 \end{split}$$

Table-1 Bias and MSE of Existing and proposed estimators

Estimators	Bias	MSE	
Existing Estimators			
Isaki[1]	5494.93	29216819.02	
Upadhyaya &Singh[2]	5483.00	29187658.80	
Kadilar & Cingi[3]	5483.00	29187658.80	
Proposed estimator	544.27	16410339.19	
M[4]			

Table-2 Percent relative efficiency of proposed estimators with existing estimators

<i>Existing</i> – <i>estimators</i> \rightarrow	Isaki	Upadhyaya & Singh Kadilar & Cingi	
Proposed – estimator	178.30	177.87	177.87

CONCLUSION

Above study clearly reveals that the proposed estimator has shown better performance in terms of bias, mean square error and percent relative efficiency as compared to other existing estimators, which can be easily seen from table-1 and table-2.

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